

Unemployment Modelled Through $M|G|\infty$ Systems

Manuel Alberto M. Ferreira^{*1}, José António Filipe^{*2}, Manuel Coelho[#]

^{*} Lisbon University Institute ISCTE-IUL, BRU-IUL
Portugal

¹manuel.ferreira@iscte.pt

²jose.filipe@iscte.pt

[#]SOCIUS, ISEG-University of Lisbon
Portugal

coelho@iseg.utl.pt

Abstract - Using results on the $M|G|\infty$ queue busy period, it is presented an application of this queue system in the unemployment periods parameters and distribution function study.

Keywords - $M|G|\infty$, busy period, unemployment.

1. The Model

In the $M|G|\infty$ queue system

- The customers arrive according to a Poisson process at rate λ
- Receive a service which time length is a positive random variable with distribution function $G(\cdot)$ and mean α
- When they arrive, each one finds immediately an available server
- Each customer service is independent from the other customers' services and from the arrivals process
- The traffic intensity is $\rho = \lambda\alpha$.

It is easy to understand how the $M|G|\infty$ queue can be applied to the unemployment study. Then

- λ is the rate at which occur the firings, supposed to occur according to a Poisson process
- The service time is the time between the worker firing and the moment he/she

finds a new job.

In a queue system a busy period is a period that begins when a customer arrives at the system finding it empty, ends when a customer abandons the system letting it empty and in it there is always at least one customer present. So in a queuing system there is a sequence of idle and busy periods.

In the $M|G|\infty$ queue system the idle periods have an exponential length with mean λ^{-1} .

But the busy period's distribution is much more complicated. In spite of it, it is possible to present some results as it will be seen.

For what interests in this work

- A busy period is a period of unemployment
- An idle period is a period of full employment.

The results that will be presented are on unemployment periods length and their number in a certain time interval.

2. Unemployment Periods Length

Call D the random variable unemployment period length. According to the results known for the $M|G|\infty$ queue busy period length distribution

$$E[D] = \frac{e^{\rho} - 1}{\lambda} \quad (2.1)$$

whichever is a worker unemployment time length distribution, see (Takács, 1962)

- As for $Var[D]$ depends on the whole unemployment time length distribution probabilistic structure. But Sathe, see (Sathe, 1985), demonstrated that

$$\lambda^{-2} \max \left[e^{2\rho} + e^\rho \rho^2 \gamma_s^2 - 2\rho e^\rho - 1; 0 \right] \leq Var[D] \leq \lambda^{-2} \left[2e^\rho (\gamma_s^2 + 1) (e^\rho - 1 - \rho) - (e^\rho - 1)^2 \right] \quad (2.2),$$

where γ_s the unemployment time length coefficient of variation

- If a worker unemployment time length distribution function is

$$G(t) = \frac{e^{-\rho}}{(1 - e^{-\rho})e^{-\lambda t} + e^{-\rho}}, t \geq 0, \quad (2.3)$$

the D distribution function is

$$D(t) = 1 - (1 - e^{-\rho})e^{-e^{-\rho}\lambda t}, t \geq 0 \quad (2.4),$$

see (Ferreira, 1991)

- If the unemployment time length of a worker is such that

$$G(t) = 1 - \frac{1}{1 - e^{-\rho} + e^{-\rho + \frac{\lambda}{1 - e^{-\rho}}t}}, t \geq 0 \quad (2.5)$$

the D distribution function is

$$D(t) = 1 - e^{-(e^\rho - 1)^{-1}\lambda t}, t \geq 0 \quad (2.6),$$

see (Ferreira, 1995)

- For α and ρ great enough (very intense unemployment conditions) since $G(\cdot)$ is such that for α great enough $G(t) \cong 0, t \geq 0$,

$$D(t) \cong 1 - e^{-\lambda e^{-\rho}t}, t \geq 0 \quad (2.7),$$

see (Ramalhoto and Ferreira, 1994).

Note:

- As for this last result, begin noting that many probability distributions fulfill the condition $G(t) \cong 0, t \geq 0$ for α great enough. The exponential distribution is one example.
- As for the meaning of α and ρ great enough, computations presented in (Ramalhoto and Ferreira, 1994) show that for $\lambda = 1$, after $\rho = 10$ it is reasonable to admit (2.7) for many distributions.

3. Mean Number of Unemployment Periods in a Time Interval

After the renewal processes theory, see (Çınlar, 1975), calling $R(t)$ the mean number of unemployment periods that begin in $[0, t]$, being $t = 0$ the beginning instant of an unemployment period, it is possible to obtain, see (Ferreira, 1995),

$$R(t) = e^{-\lambda \int_0^t [1 - G(v)] dv} + \lambda \int_0^t e^{-\lambda \int_0^u [1 - G(v)] dv} du \quad (3.1)$$

and, consequently,

$$e^{-\rho} (1 + \lambda t) \leq R(t) \leq 1 + \lambda t \quad (3.2),$$

see Ferreira (2004).

Also,

$$A) \quad G(t) = \frac{e^{-\rho}}{(1 - e^{-\rho})e^{-\lambda t} + e^{-\rho}}, t \geq 0$$

$$R(t) = 1 + \lambda e^{-\rho} t \quad (3.3)$$

$$B) \quad G(t) = 1 - \frac{1}{1 - e^{-\rho} + e^{-\rho + \frac{\lambda}{1 - e^{-\rho}}t}}, t \geq 0$$

$$R(t) = e^{-\rho} + (1 - e^{-\rho})^2 + \lambda e^{-\rho} t + e^{-\rho} (1 - e^{-\rho}) e^{\frac{\lambda}{1 - e^{-\rho}}t} \quad (3.4)$$

$$C) \quad G(t) = \begin{cases} 0, & t < \alpha \\ 1, & t \geq \alpha \end{cases}$$

$$R(t) = \begin{cases} 1, & t < \alpha \\ 1 + \lambda e^{-\rho} (t - \alpha), & t \geq \alpha \end{cases} \quad (3.5)$$

D) If the unemployment time length is exponentially distributed

$$e^{-\rho \left(1 - e^{-\frac{t}{\alpha}}\right)} + \lambda e^{-\rho} t \leq R(t) \leq e^{-\rho \left(1 - e^{-\frac{t}{\alpha}}\right)} + \lambda t \quad (3.6)$$

4. Concluding Remarks

So that this model can be applied it is necessary that the firings occur according to a Poisson process at constant rate. It is an hypothesis that must be tested.

Among the results presented, (2.1), (2.2), (2.7) and (3.2) are remarkable for its simplicity and also for requiring only the knowledge of the firings rate λ , the mean unemployment time α , and the unemployment time variance.

The other results are more complex and demand the goodness of fit test for the distributions indicated to the unemployment times.

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